# The Study of Conditional Probability Matrix for Realive Pair Genotypes 

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#### Abstract

This paper considers the conditional probability matrix for realive pair genotypes. First this paper systematically introduces HardyWeinberg equilibrium law, the joint probability distributions for relative pair genotypes and the ITO method. Then this paper discusses the calculation of probability distribution of the number of identical-by-descent allele of relative pairs, and lists the probabilities in tables. The paper gives the definition of quasi conditional probability matrix for the realive pair genotypes, and obtains a conclusion about quasi conditional probability matrix.


Keywords: quasi conditional probability matrix; ITO method; relative pair; genotype; Hardy-Weinberg equilibrium law
Mathematics Subject Classification: 62P10

## 1. Introduction

A kind of calculation method named ITO method came on (see [1]) to calculate the joint probability distribution for relative pair genotypes. With the ITO method, given the genotype of an individual, it is possible to derive the conditional probability of the genotypes of any non-inbred relative of that individual. The ITO method was extended to handle multiple alleles and was generalized for inbred populations (see [2]). The ITO method was generalized for multiple loci and was also extended to handle consanguinity (see [3]). [4] extended the ITO method to handle ordered genotypes. [5] gived an exact calculation of the probability of identity-by-descent in two-locus models using an extension of the Li-Sacks' method. Studies on the distribution of relative pair genotypes are available in the literature [6-13]. It is very important to study the conditional probability matrix for realive pair genotypes in the ITO method.

The rest of this paper is organized as follows. In Section 2, we describe Hardy-Weinberg equilibrium law. In Section 3, ITO method is introduced in
detail, and we consider the joint probability distribution matrix for relative pair genotypes. In Section 4, we discuss the calculation of probability distribution of the number of identical-by-descent allele of relative pairs. In Section 5, we give the definition of quasi conditional probability matrix for the realive pair genotypes, and obtain a conclusion about quasi conditional probability matrix. Finally, we summarize and conclude the paper in Section 6.

## 2. Hardy-Weinberg equilibrium law

British mathematician Hardy [14] and German physiologist Weinberg [15] published the equilibrium law in the genetics at the same time in 1908. HardyWeinberg equilibrium law is derived under the assumption of random mating and the principle of independent segregation. Random mating means that any woman is equally likely to marry any man. The principle of independent segregation is that a mother (or father) is equally likely to pass on either of the two alleles to her offspring (both are $1 / 2$ ), and that maternal and paternal alleles are inherited independently. Considering that there are two alleles $A$ and $a$ in a locus, it is assumed that the probabilities of the two alleles in the population of parental generation are equal to

$$
\begin{equation*}
P(A)=p, \quad P(a)=1-p=q \tag{1}
\end{equation*}
$$

If the relationship between the probabilities of three genotypes and the probabilities of alleles in a population is as follows:

$$
\begin{equation*}
P(A A)=p^{2}, \quad P(A a)=2 p q, \quad P(a a)=q^{2} \tag{2}
\end{equation*}
$$

the genotype probabilities of this population at this locus are said to have the Hardy-Weinberg proportion.

## 3. Joint probability distribution for relative pair genotypes

In order to calculate the joint probability distribution for general relative pair genotypes, a kind of mechanized calculation method named ITO method was proposed in literature [1]. ITO method is as follows. $R_{1}$ and $R_{2}$ are used to
represent genotypes of the relative pairs at the given locus respectively. Assuming that there are two alleles $A$ and $a$ at this locus, their probabilities are $p$ and $q$, respectively. If the numbers 0,1 and 2 represent three genotypes $a, A a$ and $A A$, then joint probability distribution with respect to the genotypes of relative pairs is

$$
P\left(R_{1}=i, R_{2}=j\right)=P\left(R_{1}=i \mid R_{2}=j\right) P\left(R_{2}=j\right), i, j=0,1,2
$$

where the marginal probability $P\left(R_{2}=j\right)=C_{2}^{j} p^{j}(1-p)^{2-j}$ is easily calculated by the Hardy-Weinberg equilibrium law.

The conditional probability $P\left(R_{1}=i \mid R_{2}=j\right)$ is calculated as follows. Let $I B D$ denote the number of identical-by-descent allele of relative pairs, which is a random variable with the values 0,1 and 2 . By the total probability formula,

$$
\begin{aligned}
P\left(R_{1}=i \mid R_{2}=j\right) & =\sum_{t=0}^{2} P\left(R_{1}=i \mid I B D=t, R_{2}=j\right) P\left(I B D=t \mid R_{2}=j\right) \\
& =\sum_{t=0}^{2} P\left(R_{1}=i \mid I B D=t, R_{2}=j\right) P(I B D=t)
\end{aligned}
$$

For $I B D=0,1,2$, a matrix can be used to represent the value of genotype conditional probability $p_{i j}=P\left(R_{1}=i \mid I B D=t, R_{2}=j\right)$.

When $B I D=2$, the conditional probability $p_{i j}$ is given by following matrix:

$$
\begin{align*}
& R_{1}=A A \quad A a \quad a a \quad \text { given } R_{2}= \\
& I=\left(\begin{array}{lll}
1 & 0 & 0 \\
0 & 1 & 0 \\
0 & 0 & 1
\end{array}\right) \quad \begin{array}{c}
A A \\
A a \\
a a
\end{array} \tag{3}
\end{align*}
$$

When $B I D=1$, the conditional probability $p_{i j}$ is given by following matrix:

$$
\begin{gather*}
R_{1}=A A \\
T=\left(\begin{array}{ccc}
p & q & a a \\
p / 2 & 1 / 2 & q / 2 \\
0 & p & q
\end{array}\right) \begin{array}{c}
\text { given } R_{2}= \\
A A \\
A a
\end{array} \tag{4}
\end{gather*}
$$

When $B I D=0$, the conditional probability $p_{i j}$ is given by following matrix:

$$
\begin{gather*}
R_{1}=A A \\
A a
\end{gather*} \begin{array}{lll}
a a & \text { given } R_{2}=  \tag{5}\\
O=\left(\begin{array}{lll}
p^{2} & 2 p q & q^{2} \\
p^{2} & 2 p q & q^{2} \\
p^{2} & 2 p q & q^{2}
\end{array}\right) \begin{array}{c}
A A \\
A a \\
a a
\end{array}
\end{array}
$$

Set $\Delta_{i}=P(I B D=i), i=0,1,2$. The conditional probability matrix (CPM) for relative pair genotypes is given by

$$
\begin{equation*}
W=\Delta_{2} I+\Delta_{1} T+\Delta_{0} O \tag{6}
\end{equation*}
$$

The joint probability distribution matrix for relative pair genotypes is

$$
C=\left(\begin{array}{ccc}
p^{2} & 0 & 0  \tag{7}\\
0 & 2 p q & 0 \\
0 & 0 & q^{2}
\end{array}\right) W
$$

So it's important to study $\Delta_{2}, \Delta_{1}$ and $\Delta_{0}$.

## 4. Calculation of $\Delta_{2}, \Delta_{1}$ and $\Delta_{0}$

Suppose the conditional probability matrix for the realive pair genotypes is

$$
\begin{equation*}
W=\Delta_{2} I+\Delta_{1} T+\Delta_{0} O, \Delta_{2} \neq 0 \tag{8}
\end{equation*}
$$

So let's discuss the calculation of $\Delta_{2}, \Delta_{1}$ and $\Delta_{0}$. For convenience, assume that the fathers of the realive pair have the same ancestors, and the mothers for the realive pair have the same ancestors.

Suppose the conditional probability matrix for the fathers is

$$
\begin{equation*}
W_{1}=\Delta_{21} I+\Delta_{11} T+\Delta_{01} O \tag{9}
\end{equation*}
$$

Suppose the conditional probability matrix for the mothers is

$$
\begin{equation*}
W_{2}=\Delta_{22} I+\Delta_{12} T+\Delta_{02} O \tag{10}
\end{equation*}
$$

suppose $A_{1}=$ "the realive pair share a gene identical by descent with each other, respectively from their own father".
suppose $A_{2}=$ "the realive pair share a gene identical by descent with each other, respectively from their own mother".

One has

$$
\begin{align*}
& P\left(A_{1}\right)=\frac{1}{4} \Delta_{11}+\frac{1}{2} \Delta_{21} \triangleq \Phi_{1}  \tag{11}\\
& P\left(A_{2}\right)=\frac{1}{4} \Delta_{12}+\frac{1}{2} \Delta_{22} \triangleq \Phi_{2} \tag{12}
\end{align*}
$$

In fact, $\Phi_{1}$ and $\Phi_{2}$ are both the coefficients of relationship.
Hence,

$$
\begin{aligned}
& \Delta_{2}=\Phi_{1} \Phi_{2} \\
& \Delta_{1}=\Phi_{1}\left(1-\Phi_{2}\right)+\Phi_{2}\left(1-\Phi_{1}\right) \\
& \Delta_{0}=\left(1-\Phi_{1}\right)\left(1-\Phi_{2}\right)
\end{aligned}
$$

Set vector $\Delta=\left(\Delta_{2}, \Delta_{1}, \Delta_{0}\right)$.
The first $\Delta=(1,0,0)$.
The second $\Delta$ is obtained as follows:
According to the first $\Delta=(1,0,0)$, we obtain the coefficients of relationship

$$
\Phi=\frac{1}{4} \cdot 0+\frac{1}{2} \cdot 1=\frac{1}{2} .
$$

Hence,

$$
\begin{aligned}
& \Delta_{2}=\Phi \Phi=\frac{1}{4} \\
& \Delta_{1}=\Phi(1-\Phi)+\Phi(1-\Phi)=\frac{1}{2} \\
& \Delta_{0}=(1-\Phi)(1-\Phi)=\frac{1}{4}
\end{aligned}
$$

Hence, The second $\Delta=\left(\frac{1}{4}, \frac{1}{2}, \frac{1}{4}\right)$.
The third $\Delta$ is obtained as follows:
According to the second $\Delta=\left(\frac{1}{4}, \frac{1}{2}, \frac{1}{4}\right)$, we obtain the coefficients of relationship

$$
\Phi=\frac{1}{4} \cdot \frac{1}{2}+\frac{1}{2} \cdot \frac{1}{4}=\frac{1}{4} .
$$

Hence,

$$
\begin{aligned}
& \Delta_{2}=\frac{1}{2} \cdot \frac{1}{4}=\frac{1}{8} \\
& \Delta_{1}=\frac{1}{2} \cdot\left(1-\frac{1}{4}\right)+\left(1-\frac{1}{2}\right) \cdot \frac{1}{4}=\frac{1}{8}=\frac{4}{8} \\
& \Delta_{0}=\left(1-\frac{1}{2}\right)\left(1-\frac{1}{4}\right)=\frac{3}{8}
\end{aligned}
$$

Hence, The third $\Delta=\left(\frac{1}{8}, \frac{4}{8}, \frac{3}{8}\right)$.
The rest of $\Delta \mathrm{s}$ can be done in the same manner as before.
Table 1 and Table 2 show the $\Delta \mathrm{s}$ according to the $\Phi \mathrm{s}$

$$
1 / 2,1 / 4,3 / 16,11 / 64,5 / 32,1 / 8,7 / 64,3 / 32,5 / 64,1 / 16
$$

If the conditional probability matrix for the realive pair genotypes is

$$
\begin{equation*}
W=\Delta_{1} T+\Delta_{0} O, \Delta_{1} \neq 0 \tag{13}
\end{equation*}
$$

then $\Delta_{1}=\Phi$ in Table 1 and Table 2.

## 5. Quasi conditional probability matrix for realive pair genotypes

Definition $1 W=a I+b T+c O$ is called a quasi conditional probability matrix (QCPM) for realive pair genotypes if
(1) (non-negativity) $a, b, c$ are all nonnegative rational number;
(2) (normativity) $a+b+c=1$;
(3) (decomposability) there exist real numbers $x, y$ such that $a=x y$ and $c=(1-x)(1-y)$.

In fact, (3) in Definition 1 can be replaced by the following
(3') $(b+2 a)^{2} \geq 4 a$.
The equivalence proof of (3) and (3') is as follows:
Proof: $(3) \Longrightarrow\left(3^{\prime}\right)$
From (3), we get $a=x y, c=1-(x+y)+x y$.
Hence, $x y=a, x+y=1+a-c=b+2 a$.
The real numbers $x, y$ are roots of the fillowing equation with respect to $z$,

$$
\begin{equation*}
z^{2}-(b+2 a) z+a=0 \tag{14}
\end{equation*}
$$

Obviously, $(b+2 a)^{2}-4 a \geq 0$.
$\left(3^{\prime}\right) \Longrightarrow(3)$

Since $(b+2 a)^{2} \geq 4 a$, let $x, y$ be roots of the fillowing equation with respect to $z$,

$$
\begin{equation*}
z^{2}-(b+2 a) z+a=0 \tag{15}
\end{equation*}
$$

Obviously, $x+y=b+2 a=1+a-c, x y=a$.
Hence, $a=x y, c=1-(x+y)+x y=(1-x)(1-y)$.
Obviously, CPM is a QCPM.
Lemma 1 Suppose $a+b+c=1, a \geq 0, b \geq 0, c \geq 0$. Then the following three inequalities are equivalent
(1) $4 a \leq(b+2 a)^{2}$;
(2) $(b+2 a)^{2} \leq \frac{b^{2}}{c}$;
(3) $4 a \leq \frac{b^{2}}{c}$.

Proof: $(1) \Longrightarrow(2)$
Since $(b+2 a)^{2} \geq 4 a$, let $x, y$ be roots of the fillowing equation with respect to $z$,

$$
\begin{equation*}
z^{2}-(b+2 a) z+a=0 \tag{16}
\end{equation*}
$$

so, $x+y=b+2 a, x y=a$.
Hence, $a=x y, b=x+y-2 x y, c=1-x-y+x y$.
Then

$$
\begin{aligned}
& b^{2}-c(b+2 a)^{2} \\
= & (x+y-2 x y)^{2}-(1-x-y+x y)(x+y)^{2} \\
= & {\left[(x+y)^{2}-4 x y(x+y)+4 x^{2} y^{2}\right] } \\
& -\left[(x+y)^{2}-(x+y)^{3}+x y(x+y)^{2}\right] \\
= & {\left[(x+y)^{3}-4 x y(x+y)\right] } \\
& -\left[x y(x+y)^{2}-(x+y)^{3}-4 x^{2} y^{2}\right] \\
= & (x+y)(x-y)^{2}-x y(x-y)^{2} \\
= & (x-y)^{2}(x+y-x y) \\
= & (x-y)^{2}(b+a) \geq 0,
\end{aligned}
$$

i.e.,

$$
\begin{equation*}
(b+2 a)^{2} \leq \frac{b^{2}}{c} \tag{17}
\end{equation*}
$$

$(2) \Longrightarrow(1)$
Since

$$
\begin{aligned}
& b^{2}-c(b+2 a)^{2} \\
= & b^{2}-(1-a-b)(b+2 a)^{2} \\
= & (a+b)\left[4 a(a+b)-4 a+b^{2}\right] \\
= & (a+b)\left[b^{2}+4 a b+4\left(a^{2}-a\right)\right] \geq 0,
\end{aligned}
$$

hence $b \geq-2 a+2 \sqrt{a}$, i.e.,

$$
(b+2 a)^{2} \geq 4 a
$$

$(1) \Longrightarrow(3)$
Since $(1) \Longrightarrow(2)$, so from (1) and (2), (3) holds clearly.
$(3) \Longrightarrow(1)$
Since

$$
\begin{aligned}
& b^{2}-4 a c \\
= & b^{2}-4 a(1-a-b) \\
= & b^{2}+4 a b+4\left(a^{2}-a\right) \geq 0
\end{aligned}
$$

hence $b \geq-2 a+2 \sqrt{a}$, i.e., $(b+2 a)^{2} \geq 4 a$.

Conclusion 1 If $W$ is a quasi conditional probability matrix for the realive pair genotypes, then

$$
\begin{equation*}
4 a \leq(b+2 a)^{2} \leq \frac{b^{2}}{c} \tag{18}
\end{equation*}
$$

According to Definition 1 and Lemma 1, Conclusion 1 is easy to be proved.

## 6. Conclusions

In this paper, we consider the conditional probability matrix for realive pair genotypes. We discuss the calculation of probability distribution of the number

Table 1 The first 30 sets of values of $\Delta_{2}, \Delta_{1}, \Delta_{0}, \Phi_{1}, \Phi_{2}, \Phi$

| No | $\Delta_{2}$ | $\Delta_{1}$ | $\Delta_{0}$ | $\Phi_{1}$ | $\Phi_{2}$ | $\Phi$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | 1 | 0 | 0 | 1 | 1 | $1 / 2$ |
| 2 | 1/4 | $1 / 2$ | 1/4 | $1 / 2$ | $1 / 2$ | $1 / 4$ |
| 3 | 1/8 | $1 / 2$ | 3/8 | $1 / 2$ | 1/4 | 3/16 |
| 4 | 3/32 | $1 / 2$ | 13/32 | $1 / 2$ | 3/16 | 11/64 |
| 5 | 11/128 | $1 / 2$ | 53/128 | $1 / 2$ | 11/64 | 43/256 |
| 6 | 5/64 | $1 / 2$ | 27/64 | $1 / 2$ | 5/32 | 21/128 |
| 7 | 1/16 | $1 / 2$ | 7/16 | $1 / 2$ | $1 / 8$ | 5/32 |
| 8 | 1/16 | 3/8 | 9/16 | $1 / 4$ | $1 / 4$ | $1 / 8$ |
| 9 | 7/128 | 1/2 | 57/128 | $1 / 2$ | 7/64 | 39/256 |
| 10 | 3/64 | $1 / 2$ | 29/64 | $1 / 2$ | 3/32 | 19/128 |
| 11 | 3/64 | 11/32 | 39/64 | $1 / 4$ | 3/16 | 7/64 |
| 12 | 11/256 | 43/128 | 159/256 | $1 / 4$ | 11/64 | 27/256 |
| 13 | 5/128 | $1 / 2$ | 59/128 | $1 / 2$ | 5/64 | 37/256 |
| 14 | $5 / 128$ | 21/64 | 81/128 | $1 / 4$ | 5/32 | 13/128 |
| 15 | 9/256 | 39/128 | 169/256 | 3/16 | 3/16 | 3/32 |
| 16 | $33 / 1024$ | 151/512 | 689/1024 | 3/16 | 11/64 | 23/256 |
| 17 | 1/32 | 1/2 | 15/32 | $1 / 2$ | 1/16 | 9/64 |
| 18 | 1/32 | 5/16 | 21/32 | $1 / 4$ | $1 / 8$ | 3/32 |
| 19 | 121/4096 | 583/2048 | 2809/4096 | 11/64 | 11/64 | 11/128 |
| 20 | 15/512 | 73/256 | 351/512 | 3/16 | 5/32 | 11/128 |
| 21 | 7/256 | 39/128 | 171/256 | 1/4 | 7/64 | 23/256 |
| 22 | 55/2048 | 281/1024 | 1431/2048 | 11/64 | 5/32 | 21/256 |
| 23 | 25/1024 | 135/512 | 729/1024 | 5/32 | 5/32 | 5/64 |
| 24 | 3/128 | 19/64 | 87/128 | 1/4 | 3/32 | 11/128 |
| 25 | 3/128 | 17/64 | 91/128 | 3/16 | $1 / 8$ | 5/64 |
| 26 | 11/512 | 65/256 | 371/512 | 1/8 | 11/64 | 19/256 |
| 27 | 21/1024 | 131/512 | 144/199 | $3 / 16$ | 7/64 | 19/256 |
| 28 | $5 / 256$ | 37/128 | 177/256 | $1 / 4$ | 5/64 | 21/256 |
| 29 | 5/256 | $31 / 128$ | 189/256 | 1/8 | 5/32 | 9/128 |
| 30 | 77/4096 | 499/2048 | 1703/2309 | 11/64 | 7/64 | 9/128 |

Table 2 Other sets of values of $\triangle_{2}, \Delta_{1}, \Delta_{0}, \Phi_{1}, \Phi_{2}, \Phi$

| No | $\Delta_{2}$ | $\Delta_{1}$ | $\Delta_{0}$ | $\Phi_{1}$ | $\Phi_{2}$ | $\Phi$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 31 | $9 / 512$ | $63 / 256$ | $377 / 512$ | $3 / 16$ | $3 / 32$ | $9 / 128$ |
| 32 | $35 / 2048$ | $237 / 1024$ | $1539 / 2048$ | $5 / 32$ | $7 / 64$ | $17 / 256$ |
| 33 | $33 / 2048$ | $239 / 1024$ | $1537 / 2048$ | $11 / 64$ | $3 / 32$ | $17 / 256$ |
| 34 | $1 / 64$ | $9 / 32$ | $45 / 64$ | $1 / 4$ | $1 / 16$ | $5 / 64$ |
| 35 | $1 / 64$ | $7 / 32$ | $49 / 64$ | $1 / 8$ | $1 / 8$ | $1 / 16$ |
| 36 | $15 / 1024$ | $121 / 512$ | $767 / 1024$ | $3 / 16$ | $5 / 64$ | $17 / 256$ |
| 37 | $15 / 1024$ | $113 / 512$ | $783 / 1024$ | $5 / 32$ | $3 / 32$ | $1 / 16$ |
| 38 | $7 / 512$ | $53 / 256$ | $399 / 512$ | $1 / 8$ | $7 / 64$ | $15 / 256$ |
| 39 | $55 / 4096$ | $457 / 2048$ | $3127 / 4096$ | $11 / 64$ | $5 / 64$ | $1 / 16$ |
| 40 | $25 / 2048$ | $215 / 1024$ | $1593 / 2048$ | $5 / 32$ | $5 / 64$ | $15 / 256$ |
| 41 | $49 / 4096$ | $399 / 2048$ | $3249 / 4096$ | $7 / 64$ | $7 / 64$ | $7 / 128$ |
| 42 | $3 / 256$ | $29 / 128$ | $195 / 256$ | $3 / 16$ | $1 / 16$ | $1 / 16$ |
| 43 | $3 / 256$ | $25 / 128$ | $203 / 256$ | $1 / 8$ | $3 / 32$ | $7 / 128$ |
| 44 | $11 / 1024$ | $109 / 512$ | $795 / 1024$ | $11 / 64$ | $1 / 16$ | $15 / 256$ |
| 45 | $21 / 2048$ | $187 / 1024$ | $1653 / 2048$ | $7 / 64$ | $3 / 32$ | $13 / 256$ |
| 46 | $5 / 512$ | $47 / 256$ | $413 / 512$ | $1 / 8$ | $5 / 64$ | $13 / 256$ |
| 47 | $5 / 512$ | $51 / 256$ | $405 / 512$ | $5 / 32$ | $1 / 16$ | $7 / 128$ |
| 48 | $9 / 1024$ | $87 / 512$ | $841 / 1024$ | $3 / 32$ | $3 / 32$ | $3 / 64$ |
| 49 | $35 / 4096$ | $349 / 2048$ | $3363 / 4096$ | $7 / 64$ | $5 / 64$ | $3 / 64$ |
| 55 | $5 / 1024$ | $67 / 512$ | $885 / 1024$ | $5 / 64$ | $1 / 16$ | $9 / 256$ |
| 54 | $1 / 256$ | $15 / 128$ | $225 / 256$ | $1 / 16$ | $1 / 16$ | $1 / 32$ |
| 51 | $3 / 512$ | $37 / 256$ | $435 / 512$ | $3 / 32$ | $1 / 16$ | $5 / 128$ |
| 53 | $7 / 1024$ | $81 / 512$ | $855 / 1024$ | $7 / 64$ | $1 / 16$ | $11 / 256$ |
| 53 | $25 / 4096$ | $295 / 2048$ | $3481 / 4096$ | $5 / 64$ | $5 / 64$ | $5 / 128$ |
|  | $11 / 64$ | $105 / 128$ | $1 / 8$ | $1 / 16$ | $3 / 64$ |  |
| 54 |  |  |  |  |  |  |

of identical-by-descent allele of relative pairs, and list the probabilities in tables. We give the definition of quasi conditional probability matrix for the realive pair genotypes, and obtain a conclusion about quasi conditional probability matrix.

## Reference

[1] C. C. Li, L. Sacks, The derivation of joint distribution and correlation between relatives by the use of stochastic matrices, Biometrics 10 (3) (1954) 347-360.
[2] W. H. Richardson, Frequencies of genotypes of relatives, as determined by stochastic matrices, Genetica 35 (1964) 323-354.
[3] E. R. C. Campbell M A, Relatives of probands: models for preliminary genetic analysis, Annals of human genetics 35 (1971) 225-236.
[4] D. Feng, D. E. Weeks, Ordered genotypes: An extended ito method and a general formula for genetic covariance, American Journal of Human Genetics 78 (6) (2006) 1035-1045.
[5] W. Li, An exact calculation of the probability of identity-by-descent in twolocus models using an extension of the li-sacks' method, American journal of human genetics 63 (1998) A297.
[6] M. P. Epstein, W. L. Duren, M. Boehnke, Improved inference of relationship for pairs of individuals, The American Journal of Human Genetics 67 (5) (2000) 1219-1231.
[7] R. S. Legro, R. Spielman, M. Urbanek, D. Driscoll, J. F. Strauss, A. Dunaif, Phenotype and genotype in polycystic ovary syndrome, Recent progress in hormone research 53 (1998) 217-256.
[8] D. O. Stram, C. L. Pearce, P. Bretsky, M. Freedman, J. N. Hirschhorn, D. Altshuler, L. N. Kolonel, B. E. Henderson, D. C. Thomas, Modeling and em estimation of haplotype-specific relative risks from genotype data
for a case-control study of unrelated individuals, Human heredity 55 (4) (2003) 179-190.
[9] M. Boehnke, N. J. Cox, Accurate inference of relationships in sib-pair linkage studies, The American Journal of Human Genetics 61 (2) (1997) 423429.
[10] B. A. Rybicki, R. C. Elston, The relationship between the sibling recurrence-risk ratio and genotype relative risk, The American Journal of Human Genetics 66 (2) (2000) 593-604.
[11] D. L. Harris, Genotypic covariances between inbred relatives., Genetics 50 (1964) 1319-1348.
[12] M. Crabtree, I. Tomlinson, S. Hodgson, K. Neale, R. Phillips, R. Houlston, Explaining variation in familial adenomatous polyposis: relationship between genotype and phenotype and evidence for modifier genes, Gut 51 (3) (2002) 420-423.
[13] M. R. Robinson, G. English, G. Moser, L. R. Lloyd-Jones, M. A. Triplett, Z. Zhu, I. M. Nolte, J. V. van Vliet-Ostaptchouk, H. Snieder, T. Esko, et al., Genotype-covariate interaction effects and the heritability of adult body mass index, Nature genetics 49 (8) (2017) 1174-1181.
[14] G. H. Hardy, Mendelian proportions in a mixed population, Science 28 (1954) 49-50.
[15] W. Weinberg, Über den nachweis der vererbung beim menschen, Jahreshefte des Vereins für Vaterländische Naturkunde in Württemberg 64 (1908) 368-382.

